### All of Statistics - Chapter 7 Solutions

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### 1.

Note that

$$\mathbb{E}[{\hat{F}}_n(x)] = \mathbb{E}[I(X_1 \leq x)] = \mathbb{P}(X_1 \leq x) = F(x).$$

Moreover,

$$\mathbb{V}({\hat{F}}_n(x))=\mathbb{V}(I(X_1\leq x))/n=F(x)(1-F(x))/n.$$

By the bias-variance decomposition, the MSE converges to zero. Equivalently, we can say that  $\hat{F}_n(x)$  converges to F(x) in the L2 norm. Since Lp convergence implies convergence in probability, we are done.

*Remark.* For each x,  $\hat{F}_n(x)$  is a random variable. The above proves only that each random variable  $\hat{F}_n(x)$  converges in probability to the true value of the CDF F(x). The Glivenko-Cantelli Theorem yields a much stronger result; it states that  $\|\hat{F}_n - F\|_{\infty}$  converges almost surely (and hence in probability) to zero.

#### 2.

Assumption. The Bernoulli random variables in the statement of the question are pairwise independent.

The plug-in estimator is  $\hat{p} = \overline{X}_n$ . The standard error is  $\operatorname{se}(\hat{p})^2 = \mathbb{V}(X_1)/n = p(1-p)/n$ . We can estimate the standard error by  $\widehat{\operatorname{se}}(\hat{p})^2 = \hat{p}(1-\hat{p})/n$ . By the CLT,

$$\hat{p} pprox N(p, \mathrm{se}(\hat{p})^2) pprox N(\hat{p}, \hat{\mathrm{se}}(\hat{p})^2)$$

and hence an approximate 90% confidence interval is  $\hat{p} \pm 1.64 \cdot \hat{se}(\hat{p})$ . The second part of this question is handled similarly.

#### 3.

TODO (Computer Experiment)

#### 4.

By the CLT

$$\sqrt{n}\left(rac{\sum_i I(X_i \leq x)}{n} - \mathbb{E}\left[I(X_1 \leq x)
ight]
ight) \rightsquigarrow N(0, \mathbb{V}(I(X_1 \leq x))).$$

Equivalently,

$$\sqrt{n}\left({\hat F}_n(x)-F(x)
ight) \rightsquigarrow N(0,F(x)\left(1-F(x)
ight)).$$

Or, more conveniently,

$${\hat F}_n(x)pprox N\left(F(x), rac{F(x)\left(1-F(x)
ight)}{n}
ight).$$

*Remark.* The closer (respectively, further) F(x) is to 0.5, the more (respectively, less) variance there is in the empirical distribution evaluated at x.

#### 5.

Without loss of generality, assume x < y. Then,

$$egin{aligned} ext{Cov}({\hat F}_n(x),{\hat F}_n(y)) &= rac{1}{n^2} ext{Cov}(\sum_i I(X_i \leq x),\sum_i I(X_i \leq y)) \ &= rac{1}{n^2}\sum_i ext{Cov}(I(X_i \leq x),I(X_i \leq y)) = rac{1}{n} ext{Cov}(I(X_1 \leq x),I(X_1 \leq y)) \ &= rac{1}{n}(F(x)-F(x)F(y)) = rac{1}{n}F(x)\left(1-F(y)
ight). \end{aligned}$$

#### **6**.

By the results of the previous question,

$$egin{aligned} n \cdot ext{se}(\hat{ heta})^2 &= n \mathbb{V}(\hat{F}_n(b) - \hat{F}_n(a)) \ &= n \mathbb{V}(\hat{F}_n(b)) + n \mathbb{V}(\hat{F}_n(a)) - 2n \operatorname{Cov}(\hat{F}_n(b), \hat{F}_n(a)) \ &= F(b) \left(1 - F(b)\right) + F(a) \left(1 - F(a)\right) - 2F(a) \left(1 - F(b)
ight) \ &= \left(F(b) - F(a)
ight) \left[1 - \left(F(b) - F(a)
ight)
ight]. \end{aligned}$$

We can use the estimator

$$\hat{ ext{se}}(\hat{ heta})^2 = rac{1}{n} \Big( {\hat{F}}_n(b) - {\hat{F}}_n(a) \Big) \left[ 1 - \left( {\hat{F}}_n(b) - {\hat{F}}_n(a) 
ight) 
ight].$$

An approximate 1-lpha confidence interval is  $\hat{ heta}\pm z_{lpha/2}\cdot\hat{
m se}(\hat{ heta}).$ 

*Remark*. The closer F(b) - F(a) is to zero or one, the smaller the standard error.

#### 7.

TODO (Computer Experiment)

#### 8.

# 9.

This is an application of our findings in Question 2. In particular, we use the estimate (90 - 85)/100 = 0.05. A  $1 - \alpha$  confidence interval for this estimate is  $0.05 \pm z_{\alpha/2} \cdot \hat{se}$  where

$$\hat{
m se} = \sqrt{0.9 \left(1 - 0.9
ight) / 100 + 0.85 \left(1 - 0.85
ight) / 100} pprox 0.047.$$

The z-scores corresponding to 80% and 95% intervals are approximately 1.28 and 1.96.

## 10.

TODO (Computer Experiment)